

## Interval Estimation

Underlying distribution is Normal—Central Limit theorem

Example: Suppose we draw a sample of children and measure Diastolic Blood Pressure ( $n=10$ )

$$\bar{x} = 58 \text{ mm Hg}$$

Not certain that  $\mu = 58$  exactly—second sample estimate might be  $\bar{x} = 61$  mm Hg

Assume DBP  $\sim N(\mu, \sigma^2)$

Therefore, if  $\mu, \sigma^2$  were known, and because

$$\bar{x} \sim N(\mu, \sigma^2 / n)$$

95% of all sample means would fall within

$$\left( \mu - 1.96 \frac{\sigma}{\sqrt{n}}, \mu + 1.96 \frac{\sigma}{\sqrt{n}} \right)$$

Similarly, if we convert  $\bar{x}$  to  $z$  where

$$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}, \text{ then } z \sim N(0, 1)$$

Therefore, 95% of the  $z$  values are between (-1.96, 1.96)

**$t$  distribution**

Calculate  $t = \frac{\bar{x} - \mu}{s / \sqrt{n}}$

Shape depends on  $n$ —not a unique distribution but a family indexed by a parameter called degrees of freedom

If  $X_1, X_2, \dots, X_n \sim N(\mu, \sigma^2)$  are independent,

$$t = \frac{\bar{X} - \mu}{s / \sqrt{n}} \sim t_{n-1}$$

Notation:  $t_{d,u}$

Where  $d$  = degrees of freedom

$u$  refers to some percentile, like 95<sup>th</sup>

$$\Pr(t_d < t_{d,u}) = u$$

For example:  $t_{9, .95}$

### Properties of $t$ distribution

1. Symmetric about 0
2. More spread out than  $N(0, 1)$
3. For any percentile,  $\alpha$ ,  $\alpha > 0.5$ ,

$t_{d, 1-\alpha}$  is always larger than corresponding percentile for  $N(0, 1)$ — $z_{1-\alpha}$

4. As  $d$  increases,  $t$  approaches  $N(0, 1)$

5.  $t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$  has more variability than

$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

### Using the $t$ -table (Page 823)

$$t_{10, 0.975} =$$

$$t_{20, 0.975} =$$

$$t_{40, 0.975} =$$

$$t_{120, 0.975} =$$

$$z_{0.975} =$$

$$t_{20, 0.025} =$$

$$t_{20, 0.05} =$$

## Confidence Intervals

$t$  statistic follows a  $t$  distribution with  $n-1$  degrees of freedom

$$\Pr(t_{n-1, .025} < t < t_{n-1, .975}) = 0.95$$

or

$$\Pr(t_{n-1, \frac{\alpha}{2}} < t < t_{n-1, 1-\frac{\alpha}{2}}) = 1 - \alpha$$

that is,

$$t_{n-1, \frac{\alpha}{2}} < \frac{\bar{x} - \mu}{s/\sqrt{n}} \quad \text{and} \quad \frac{\bar{x} - \mu}{s/\sqrt{n}} < t_{n-1, 1-\frac{\alpha}{2}}$$

therefore,

$$\Pr(\bar{x} - t_{n-1, 1-\frac{\alpha}{2}} * s/\sqrt{n} < \mu < \bar{x} + t_{n-1, 1-\frac{\alpha}{2}} * s/\sqrt{n}) = 1 - \alpha$$

therefore,

$$(\bar{x} - t_{n-1, 1-\frac{\alpha}{2}} * s/\sqrt{n}, \bar{x} + t_{n-1, 1-\frac{\alpha}{2}} * s/\sqrt{n}) = 100\%(1 - \alpha) \text{ CI}$$

or

$$(\bar{x} \pm t_{n-1, 1-\frac{\alpha}{2}} * s/\sqrt{n}) = 100\%(1 - \alpha) \text{ CI}$$

Example: if  $n = 12$ , degrees of freedom = 11

$$95\% \text{ CI} = (\bar{x} - 2.201 * s/\sqrt{12}, \bar{x} + 2.201 * s/\sqrt{12})$$

Example: Estimate the average level of a certain enzyme in a human population. Take a sample of 10 individuals and determine the level of the enzyme in each.

$$\bar{x} = 22$$

$$s^2 = 45$$

$$\begin{aligned} 95\% \text{ CI} &= (22 - 2.262 * \frac{\sqrt{45}}{\sqrt{10}}, 22 + 2.262 * \frac{\sqrt{45}}{\sqrt{10}}) \\ &= (22 - 2.262 * 2.1213, 22 + 2.262 * 2.1213) \\ &= (22 - 4.798, 22 + 4.798) \\ &= (17.202, 26.798) \end{aligned}$$

For  $\sigma$  known, use