

## Inference for the Mean of a Population

Outline:

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- The Matched Pairs  $t$ -test.

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### Student's $t$ Distribution

If the true population SD,  $\sigma$ , is unknown, we estimate  $\sigma$  using the sample SD  $S$ .

When the standard deviation of a statistic is estimated from the data, the result is called the standard error of the statistic. The standard error of the sample mean  $\bar{x}$  is  $SE_{\bar{X}} = \frac{S}{\sqrt{n}}$ .

Now, instead of dealing with

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

we are interested in the quantity

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{(n-1)}$$

Here,  $t_{(n-1)}$  is **Student's  $t$  distribution**, with  $n - 1$  *degrees of freedom*. The table for the  $t$  distribution is given in the back of the book.

(Unlike the Normal or Binomial distributions, each of which has two parameters, the  $t$  distribution has only one parameter, called the **degrees of freedom**.)

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### Inference for the Mean of a Population

We have already seen that when we take a SRS,  $x_1, x_2, \dots, x_n$ , from a population with unknown  $\mu$  and known  $\sigma$ , if either

- the population is Normally distributed, or
- $n$  is large enough,

then

$$\bar{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$

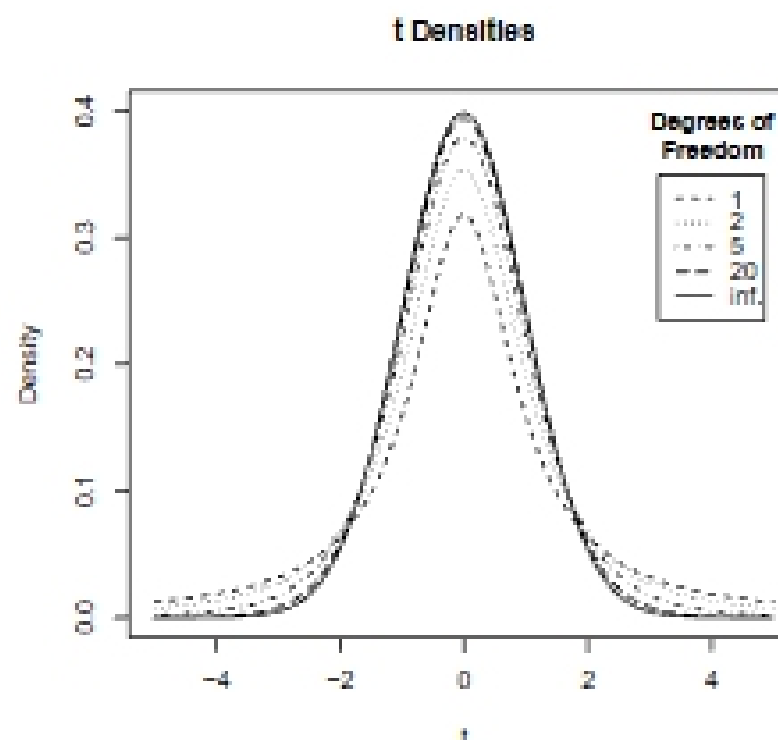
Equivalently, the standardized sample mean, or the one-sample  $z$  statistic is

$$z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

What if the population SD  $\sigma$  is unknown? (A far more likely occurrence.)

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### The Density of Student's $t$



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### Properties of the $t$ Distribution

- Symmetric about zero
- Bell-shaped – similar to normal distribution
- More spread out than normal – heavier tails
- Exact shape depends on the degrees of freedom
- As the number of degrees of freedom increases, the  $t$  distribution converges to the Normal distribution.

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For this question,  $\alpha = .05$ . For a  $t_{30}$  density, the area to the right of 2.042 is  $\frac{.05}{2} = .025$  and the area to the left of  $-2.042$  is  $.025$ , so  $t^* = 2.042$

For a 95% confidence interval for a  $t$  distribution with 100 degrees of freedom,  $t^* = 1.984$ .

For a 95% confidence interval for a  $t$  distribution with 1000 degrees of freedom,  $t^* = 1.962$ .

Note that a  $z^* = 1.960$  for a 95% confidence interval for a standard normal distribution.

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### Confidence Intervals with Unknown $\sigma$

Recall that, for a population with unknown  $\mu$  and known  $\sigma$ ,  $100(1 - \alpha)\%$  CI for  $\mu$ , based on an SRS  $x_1, x_2, \dots, x_n$  is given by

$$\left( \bar{x} - z^* \frac{\sigma}{\sqrt{n}}, \bar{x} + z^* \frac{\sigma}{\sqrt{n}} \right)$$

If  $\sigma$  is unknown,

$$\frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{(n-1)}$$

so we substitute a  $t$ -critical value,  $t^*$  for  $z^*$ :

$$\left( \bar{x} - t^* \frac{s}{\sqrt{n}}, \bar{x} + t^* \frac{s}{\sqrt{n}} \right)$$

The critical value,  $t^*$ , is chosen such that  $100(1 - \alpha)\%$  of the area under the  $t_{(n-1)}$  density lies between  $-t^*$  and  $t^*$ .

The area to the right of  $t^*$  should be  $\frac{\alpha}{2}$  and the area to the left of  $-t^*$  should be  $\frac{\alpha}{2}$ .

What would  $t^*$  be for a 95% confidence interval for a  $t$  distribution with 30 degrees of freedom, i.e.  $n = 31$ ?

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If the underlying population is Normally distributed, the confidence interval

$$\left( \bar{x} - t^* \frac{s}{\sqrt{n}}, \bar{x} + t^* \frac{s}{\sqrt{n}} \right)$$

is exact.

Otherwise, if  $n$  is not too small ( $n \geq 15$ ), the data are not strongly skewed, and there are no outliers, the interval is approximately correct.

With  $n$  sufficiently large (say  $n \geq 40$ ), the approximation is correct even if the data are clearly skewed.

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### A One-sample $t$ -test

Suppose a SRS of size  $n$  is drawn from a  $N(\mu, \sigma)$  population with both  $\mu$  and  $\sigma$  unknown. The  $t$ -statistic,

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}}$$

has the  $t$  distribution with  $n - 1$  d.f.

To test  $H_0 : \mu = \mu_0$ , compute the one-sample  $t$  statistic,

$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$$

The  $p$ -values are

$$H_a : \mu > \mu_0 \quad P(t_{n-1} \geq t)$$

$$H_a : \mu < \mu_0 \quad P(t_{n-1} \leq t)$$

$$H_a : \mu \neq \mu_0 \quad 2P(t_{n-1} \geq |t|)$$

These are exact if the population is normal, and otherwise approximately correct for large  $n$ .

There is evidence that the population mean tumor growth is not  $4mm$ .

What if we wanted a 99% CI for  $\mu$  instead? For this case  $\alpha = .01$  We need to first need to find  $t^*$  for a  $t$  distribution with 19 degrees of freedom.

What is  $t^*$ ?

The CI is given by

$$\begin{aligned} & \left( \bar{x} - t^* \frac{s}{\sqrt{n}}, \bar{x} + t^* \frac{s}{\sqrt{n}} \right) = \\ & = \left( 3.8 - 2.861 \times \frac{0.3}{\sqrt{20}}, 3.8 + 2.861 \times \frac{0.3}{\sqrt{20}} \right) = \\ & = (3.61, 3.99) \end{aligned}$$

where  $t^* = 2.861$  is the upper 0.005 critical value of  $t_{19}$ .

Note that 4 is outside this CI. From this, we can draw the same conclusion as from the test.

Namely, at significance level  $\alpha = 0.01$ , the mean growth not equal to  $4mm$ .

### Example

Let  $X$  (in  $mm$ ) denote the growth in 15 days of a tumor induced in a mouse. It is known from a previous experiment that the average tumor growth is  $4mm$ . A sample of 20 mice that have a genetic variant hypothesized to be involved in tumor growth yielded  $\bar{x} = 3.8mm$ ,  $s = 0.3mm$ . Test whether  $\mu = 4$  or not, assuming growths are normally distributed.

1. State the hypotheses

$$H_0 : \mu = 4 \quad H_a : \mu \neq 4$$

2. Calculate the  $t$ -statistic

$$t = \frac{3.8 - 4.0}{0.3/\sqrt{20}} = -2.98$$

3. Determine the  $p$ -value using the  $t$ -distribution table in the book

$$p = 2P(t_{19} \geq 2.98) < 2P(t_{19} \geq 2.861) = 2(.005) = .01$$

Since  $p$  is less than 0.01, we reject  $H_0$  at significance level  $\alpha = 0.01$ . ( $p$ -value=0.008)

### Confidence Intervals and Two-Sided Tests

A two-sided hypothesis test with significance level  $\alpha$  rejects the null hypothesis  $H_0 : \mu = \mu_0$  if and only if the value  $\mu_0$  falls outside the  $100(1 - \alpha)\%$  CI for  $\mu$ .

Reporting a CI is generally more informative than just reporting a  $p$ -value or the decision made on the basis of a hypothesis test.

### One-Sided Alternatives

In the previous example, suppose we wished to test whether  $\mu < 4$ .

1. State the hypotheses

$$H_0 : \mu = 4 \quad H_a : \mu < 4$$

2. Calculate the  $t$ -statistic

$$t = \frac{3.8 - 4}{0.3/\sqrt{20}} = -2.98$$